

# Reliability and accuracy of heart rate variability metrics versus ECG segment duration

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**Abstract** Despite the exponential growth in heart rate variability (HRV) research, the reproducibility and reliability of HRV metrics continues to be debated. We estimated the reliability of 11 metrics calculated from 5 min records. We also compared the accuracy of the HRV metrics calculated from ECG records spanning 10 s to 10 min as compared with the metrics calculated from 5 min records. The mean heart rate was more reproducible and could be more accurately estimated from very short segments (<1 min) than any of the other HRV metrics. HRV metrics that effectively highpass filter the R–R interval series were more reliable than the other metrics and could be more accurately estimated from very short segments. This indicates that most of the HRV is caused by drift and nonstationary effects. Metrics that are sensitive to low frequency components of HRV have poor repeatability and cannot be estimated accurately from short segments (<10 min).

**Keywords** Heart rate variability (HRV) · Interbeat intervals (IBI) · Intraclass correlation

coefficient (ICC) · Reliability · Reproducibility · Repeatability · Consistency

## 1 Introduction

Heart rate variability (HRV) has received a tremendous amount of attention since the seminal work of Akselrod et al. [1]. Established clinical applications of HRV include risk assessment of patients after myocardial infarction and early diagnosis of diabetic autonomic neuropathy [24, 57]. The physiologic mechanisms underlying HRV continue to be investigated [7, 26, 36, 52].

A task force of the European Society of Cardiology and the North American Society of Pacing and Electrophysiology published standards of measurement, interpretation and use of HRV in 1996 [57]. The task force specified many different HRV metrics for both short-term records (5 min) and long-term records (24 h). Although many other measures of HRV have been proposed and investigated, those specified by the task force have been the most widely applied.

Despite the exponential growth in HRV research and the number of studies published addressing the reliability, repeatability and reproducibility of HRV metrics, this topic continues to be debated. Some studies concluded that their examined HRV metrics were reliable [2, 5, 8, 13, 14, 18, 19, 25, 29, 31, 34, 35, 39, 41, 50, 55, 58, 59, 60, 61], while others found the reliability to be moderate or low [6, 23, 33, 51, 56, 62]. This discrepancy of results may be due, in part, to the fact that these studies examined distinct subject populations, and their conclusions on HRV reliability apply only to the specific HRV metrics studied.

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In this paper, we address how the accuracy of time- and frequency-domain HRV metrics varies with the record duration and estimate the reliability of these metrics calculated from 5 min records. In order to address the question of HRV reliability, we studied the reliability of the most widely used HRV metrics in three different databases. We used publicly available databases to ensure the reproducibility of our results. Additionally, we propose a new methodology and visualization technique which can be used to assess the segment length required for a given degree of accuracy.

## 2 HRV metrics

We studied four time-domain and six frequency-domain metrics, recommended by the task force for short-term ECG records. We also studied approximate entropy as a measure of the complexity of HRV. We denote each of the  $N$  beat times as  $t(n)$  for  $n \in \{1, \dots, N\}$ . We denote the interval between beats as  $\delta(n) = t(n) - t(n-1)$ . As recommended by the task force, we define the time of occurrence of each interval  $\delta(n)$  as  $t(n)$ .

### 2.1 Time domain metrics

#### 2.1.1 Standard deviation of the normal-to-normal

We calculated the standard deviation of the normal-to-normal (NN) intervals (SDNN) as

$$\text{SDNN} = \sqrt{\frac{1}{N-2} \sum_{n=2}^N (\delta(n) - \bar{\delta})^2} \quad (1)$$

where  $\bar{\delta}$  is the average NN interval,

$$\bar{\delta} = \frac{1}{N-1} \sum_{n=2}^N \delta(n). \quad (2)$$

The scaling factor is  $N-2$  because there are  $N-1$  intervals in the record and one degree of freedom is used to estimate the mean NN interval.

#### 2.1.2 HRV triangular index

The HRV triangular index (HRVTI) is a measure of the shape of the NN interval distribution. Generally, uniform distributions representing large variability have large values and distributions with single large peaks have small values. The metric is defined in terms of a histogram of the NN intervals. Here, we represent

the number of intervals in the  $i$ th bin centered at  $t_i$  as  $b(t_i)$ . HRVTI is then defined as

$$\text{HRVTI} = \frac{\sum_{i=1}^{N_b} b(t_i)}{\max_i b(t_i)} = \frac{N-1}{\max_i b(t_i)} \quad (3)$$

where  $N_b$  is the number of bins. We used a constant bin width of  $1/(f_s) = 8$  ms. The task force did not specify how to align the bins. We chose to locate the bin centers at integer multiples of  $1/f_s$  so that each bin only contained equal intervals.

#### 2.1.3 Root mean square of successive NN interval differences

We calculated the root mean square of successive NN interval differences (RMSSD) as

$$\text{RMSSD} = \sqrt{\frac{1}{N-2} \sum_{n=3}^N [\delta(n) - \delta(n-1)]^2}. \quad (4)$$

#### 2.1.4 Standard deviation of successive NN interval differences

We calculated the standard deviation of successive NN interval differences (SDSD) as

$$\text{SDSD} = \sqrt{\frac{1}{N-3} \sum_{n=3}^N [\delta(n) - \delta(n-1) - \bar{\delta}_\delta]^2} \quad (5)$$

where  $\bar{\delta}_\delta$  is given by

$$\bar{\delta}_\delta = \frac{1}{N-2} \sum_{n=3}^N \delta(n) - \delta(n-1). \quad (6)$$

In most cases  $\bar{\delta}_\delta$  is nearly zero and there is little difference between RMSSD and SDSD.

### 2.2 Frequency domain metrics

All frequency domain HRV metrics are based on the estimated power spectral density (PSD) of the NN intervals. Although the task force gave specific definitions of the metrics, it did not specify how to estimate the PSD. There are many methods of estimating PSD and each generates different HRV metric values. If the NN interval series is not well behaved or too short, as in our case, these differences can be substantial. In this section we give a complete description of our PSD estimator, as required by the task force.

We smoothed and uniformly resampled the NN intervals at a rate of 3 Hz with a kernel smoother,

$$d(k) = \frac{\sum_{n=1}^N \delta(n) b\left(\frac{|kT_s - t(n)|}{\sigma_b}\right)}{\sum_{n=1}^N b\left(\frac{|kT_s - t(n)|}{\sigma_b}\right)} \quad (7)$$

where  $T_s = 1/3$  s is the resample interval,  $\sigma_b$  is the kernel width, and  $b(\cdot)$  is a clipped and scaled Gaussian kernel function

$$\begin{cases} \exp(-u^2/2) & -5 \leq u \leq 5 \\ 0 & \text{otherwise.} \end{cases} \quad (8)$$

We generated the uniformly resampled signal  $d(k)$  over the duration of the ECG record. The kernel smoother also acts like a lowpass filter with a cutoff frequency determined by  $\sigma_b$ . Large values of  $\sigma_b$  reduce aliasing but can bias HRV metrics that rely on accurate estimates of the PSD at higher frequencies. We used  $\sigma_b = 0.25$  s.

We applied Welch’s nonparametric method of Periodogram averaging to estimate the PSD of  $d(k)$ . This estimate is calculated as the average of Periodograms calculated from overlapping segments of  $d(k)$ ,

$$\hat{R}(e^{j\omega}) = \frac{1}{KL} \sum_{i=0}^{K-1} \left| \sum_{k=0}^{L-1} (d(k+iL) - \bar{d}) w(k) e^{-j\omega k} \right|^2 \quad (9)$$

where  $K$  is the number of segments,  $L$  is the number of samples in each segment,  $\bar{d}$  is the sample average of  $d(k)$  over the full record, and  $w(k)$  is a window function that determines the tradeoff between main lobe width and sideband leakage. We used a Blackman window with a length of 20 s or the duration of the record, whichever is shorter. We allowed a 50% overlap between segments. Each segment was padded with zeros to a total length of 4,096 samples to minimize the error in estimating the signal power over specified frequency ranges with Riemann sums.

### 2.2.1 Low frequency

The low-frequency (LF) power was calculated as the total signal power in the frequency range of 0.04–0.15 Hz.

### 2.2.2 High frequency

The high-frequency (HF) power was calculated as the total signal power in the frequency range of 0.15–0.40 Hz.

### 2.2.3 LF norm

The low-frequency normalized (LF norm) power was calculated as

$$\text{LF norm} = 100 \times \frac{\text{LF}}{\text{TP} - \text{VLF}} \quad (10)$$

where TP is the total signal power is defined below and VLF is defined as the total signal power at frequencies less than 0.04 Hz.

### 2.2.4 HF norm

The high frequency normalized (HF Norm) power was calculated

$$\text{HF norm} = 100 \times \frac{\text{HF}}{\text{TP} - \text{VLF}}. \quad (11)$$

### 2.2.5 Low frequency/high frequency

The low frequency–high frequency ratio (LF/HF) was calculated as LF/HF.

### 2.2.6 Total power

The total power (TP) was calculated as the integral of the PSD estimate over the full frequency range of 0.0–1.5 Hz. This is approximately equal to the variance of  $d(k)$ .

## 2.3 Complexity metrics

### 2.3.1 Approximate entropy

Approximate entropy is one of the nonlinear metrics which has been used to analyze the  $R$ – $R$  intervals in order to estimate reductions of complexity associated with specific pathologies [32, 42]. Approximate entropy was introduced as a quantification of regularity in sequences and time series data, initially motivated by applications to relatively short, noisy data sets. Mathematically it is part of a general development of approximating Markov Chains to a process [44]. It provides a finite sequence formulation of randomness, via proximity to maximal irregularity [47, 49]. A statistical evaluation of approximate entropy (ApEn) is available in [46]. ApEn is scale invariant and model independent, evaluates both dominant and subordinate patterns in data, and discriminates series for which clear feature recognition is difficult. It measures the

logarithmic likelihood that runs of patterns that are close remain close on subsequent incremental comparisons, and assigns a non-negative number to a time series, with larger values corresponding to more complexity or irregularity in the data. ApEn has two user-specified parameters: a run length  $m$  and a tolerance window  $r$ . It is important to consider ApEn( $m, r$ ), or ApEn( $m, r, N$ ), where  $N$  is the number of points of the time series, as a family of parameters: comparisons between time series segments can only be made with the same values of  $m$  and  $r$  [45]. Formally, given  $N$  data points from a time series  $\{x(1), x(2), \dots, x(N)\}$ , the ApEn is computed according to:

- From a sequence of  $m$ -vectors  $X(1), X(2), \dots, X(N - m + 1)$  defined as  $X(i) = [x(i), x(i + 1), \dots, x(i + m - 1)]$ ,  $i = 1, 2, \dots, N - m + 1$ . These vectors represent  $m$  consecutive  $x$  values, commencing with the  $i$ th point.
- Calculate the distance between  $X(i)$  and  $X(j)$ ,  $d[X(i), X(j)]$ , as the maximum absolute difference between their respective scalar components,

$$d[X(i), X(j)] = \max_{k=0,1,\dots,m-1} (|x(i+k) - x(j+k)|) \quad (12)$$

- For a given  $X(i)$ , count the number of  $j$  ( $j = 1, 2, \dots, N - m + 1$ ) for  $j \neq i$  such that  $d = [X(i), X(j)] \leq r$ , denoted as  $N^m(i)$ . Then, for  $i = 1, 2, \dots, N - m + 1$ ,

$$C_r^m(i) = \frac{N^m(i)}{N - m + 1} \quad (13)$$

The  $C_r^m(i)$  values measure, within a tolerance  $r$ , the regularity (frequency) of patterns similar to a given segment of length  $m$ .

- Compute the natural logarithm of each  $C_r^m(i)$ , and compute the average of it over  $i$ ,

$$\phi^m(r) = \frac{1}{N - m + 1} \sum_{i=1}^{N-m+1} \ln C_r^m(i) \quad (14)$$

where  $\phi^m(r)$  represents the average frequency of all the  $m$ -point patterns in the sequence remain close to each other.

- Increase the dimension to  $m + 1$ . Repeat steps (1) to (4) and find  $C_r^{m+1}$  and  $\phi^{m+1}(r)$ .
- Theoretically, the ApEn is defined as,

$$\text{ApEn}(m, r) = \lim_{N \rightarrow \infty} [\phi^m(r) - \phi^{m+1}(r)]. \quad (15)$$

There are two ways to look at ApEn. From one point of view, it is a statistical metric (the average of

the logarithm of a conditional probability), which makes it applicable to both deterministic and stochastic processes. From another point of view, it reflects the rate of new pattern generation and is thus related to the concept of entropy [22].

In practice, the number of data points  $N$  is finite. We implemented this formula by defining the statistic [43]:

$$\text{ApEn}(m, r) = \phi^m(r) - \phi^{m+1}(r). \quad (16)$$

In this study, we calculated the estimated ApEn of the interbeat interval series  $\delta(n)$ . ApEn was estimated with the established parameter values of  $m = 1$  and  $r = 0.25 \sigma_\delta$ , where  $\sigma_\delta$  is the sample standard deviation of  $\delta(n)$ . Normalizing  $r$  in this manner gives ApEn a translation and scale invariance, in that it remains unchanged under uniform process magnification, reduction, or constant shift to higher or lower values [45]. Several studies have demonstrated that these input parameters produce good statistical reproducibility for ApEn [27, 43, 48].

### 3 Duration/reliability analysis

#### 3.1 Duration analysis

##### 3.1.1 Patient population

We used the normal sinus rhythm R–R interval database posted on PhysioNet [40]. The database includes the beat times of 54 long-term ambulatory ECG recordings (21.4–24.2 h, 1,280 h total). The recordings were acquired from 30 men (ages 28.5–76 years) and 24 women (ages 58–73 years). The original recordings were sampled at a rate of 128 Hz. The beat and rhythm annotations were generated by an automatic algorithm and then manually reviewed and corrected.

##### 3.1.2 Record sampling

We randomly selected 1,000 15-min records with replacement from the database. For each case, we randomly selected a subject and a 15 min record from the entire recording. We then discarded all records did not consist entirely of normal sinus rhythm beats. This reduced the number of usable records to 696.

##### 3.1.3 Figure of merit

We calculated each of the 11 HRV metrics using segments of each 15 min ECG record. The segment durations ranged from 10 s to 10 min. We treated a

5 min segment in the center of each 15 min record (300–600 s) as a short-term ECG baseline record. We randomly selected segments of 5 min or less from the 5 min baseline. We selected segments of more than 5 min randomly from the 15 min record, but the selection was constrained to include all of the 5 min baseline. For each duration, we calculated the percent difference between the HRV metric calculated from the randomly selected segment and the 5 min baseline,

$$\epsilon(i, \tau) = 100 \times \frac{m(i, \tau) - m_{BL}(i)}{\sigma_{BL}} \quad (17)$$

where  $\epsilon(i, \tau)$  is the percent difference,  $i \in \{1, \dots, 696\}$  is the record index,  $\tau$  is the duration of the randomly selected segment,  $m(i, \tau)$  is the metric calculated from the randomly selected segment,  $m_{BL}(i)$  is the metric calculated from the baseline, and  $\sigma_{BL}$  is the sample standard deviation of the metric calculated from the baselines of all 696 records.

### 3.2 Reliability analysis

#### 3.2.1 Patient population

We used three data bases posted on PhysioNet [40]. The databases included the normal sinus rhythm RR interval database described previously, the European ST-T database, and the cardiac arrhythmia suppression trail (CAST) database. The European ST-T database consists of 90 annotated ambulatory recordings from 79 subjects. Each subject was diagnosed or suspected of having myocardial ischemia. The database includes patients with hypertension, ventricular dyskinesia. Some subjects were taking medication that may have affected their recordings. The original signals were sampled at 250 Hz.

The CAST database consisted of 762 pre-treatment recordings from patients who had an acute myocardial infarction within the preceding 2 years. Most of the signals were sampled at 128 Hz. All others were sampled at 125.4 Hz.

All three databases contained annotations that were generated by an automatic algorithm and then manually reviewed and corrected by one or more experts.

#### 3.2.2 Record sampling

For each recording, we randomly selected a 15-min segment. If the segment did not consist entirely of normal sinus rhythm beats, we resampled the recording. Up to 500 attempts were made to find a 15-min segment with normal sinus rhythm beats. This resulted

in fifty 15-min records out of 54 possible for the Normal Sinus Rhythm database, 41 out of 90 possible for the European ST-T database, and 480 out of 762 possible for the CAST database.

#### 3.2.3 Intraclass correlation coefficients

Reliability analysis is a means of assessing what fraction of the variability in observed values is due to measurement error. We used intraclass correlation coefficients (ICC) as our measure of reliability [38, 53]. The ICC is defined as a ratio of explained variation to total variation. The calculation of these coefficients requires repeated measurements of the same entity, but allows for random errors in each measurement. There are several types of ICC that one may choose from depending on the statistical model of the study. We chose a one-way random effects model that can be estimated with an analysis of variance (ANOVA) model II with random factor levels. This treats each observation as a realization of the following random process

$$y_{ij} = \mu_i + \epsilon_{ij} \quad (18)$$

where  $i$  is the record index,  $j$  is the measurement index,  $\mu_i$  are independent and normally distributed  $N(\mu, \sigma_\mu^2)$ ,  $\epsilon_{ij}$  are independent and normally distributed  $N(0, \sigma^2)$ , and  $\mu_i$  and  $\epsilon_{ij}$  are mutually independent.

We treated the metrics calculated from the first (0–300 s) and last (600–900 s) 5 min segments of each 15 min record as a repeated measurement of the same entity. The ICC was calculated as,

$$\hat{\rho} = \frac{\text{MSTR} - \text{MSE}}{\text{MSTR} + (M - 1)\text{MSE}} = \frac{\hat{\sigma}_\mu^2}{\hat{\sigma}_\mu^2 + \hat{\sigma}^2} \quad (19)$$

where  $M = 2$  is the number of measurements per entity, MSTR is the treatment mean square, and MSE is the error mean square. This is an estimate of the true ICC defined as

$$\rho = \frac{\sigma_\mu^2}{\sigma_\mu^2 + \sigma^2}. \quad (20)$$

The estimate is biased, but consistent and tolerant of modest departures from normality [53]. Values close to one indicate that the variation between entities is significantly greater than measurement error and values close to zero indicate that the measurement error is too large to accurately discern differences between entities.

## 4 Results and discussion

Our results indicate that all the HRV metrics included in this study are sensitive to changes in duration of the ECG segment (Fig. 1). In these plots we show the median and four ranges of the percent difference in (17) for ECG segment durations ranging from 10 s to 10 min. From these results we also conclude that most of the HRV metrics are biased estimates. This confirms the task force's statement that metrics calculated from segments of different durations are not comparable [57].

The plots in Fig. 1 indicate that all the HRV metrics are sensitive to changes in the duration of the ECG segment. Most of the total HRV is caused by drift and nonstationary effects that cannot be estimated accurately from short segments. HF, SDDSD, and RMSSD had the best overall performance because they both effectively highpass filter the NN interval series and are thereby less sensitive to these effects than the other metrics. The bias and variance of HF, SDDSD, and RMSSD were small and comparable to the mean heart rate (MHR).

The results of the reliability study (Tables 1, 2, 3) demonstrate that the MHR was much more reproducible than any of the other HRV metrics. RMSSD, SDDSD, and HF were significantly more reliable than the other metrics and achieved an ICC comparable to MHR. ApEn was the most unreliable metric of the HRV measures tested in this study. We also note that pairs of 5 min segments spaced 5 min apart are not actually repeated measurements of the same entity since the NN interval series is known to be a nonstationary process. Here we assumed that the process is locally stationary and that the metrics calculated from each pair of segments can be treated as independent samples drawn from the same distribution.

It is important to recognize that reliability depends on the variability between subjects, that is, the second column of Tables 1, 2, and 3 varies across databases. This may explain the discrepancy in the ICCs across the databases, and the discrepancy of the results of reliability studies on HRV metrics reported in the literature. However, the first column should be reasonably consistent across databases. Additionally, the precise measurement of the NN intervals also affect the reliability of the HRV metrics [4, 10, 11, 12, 15, 16].

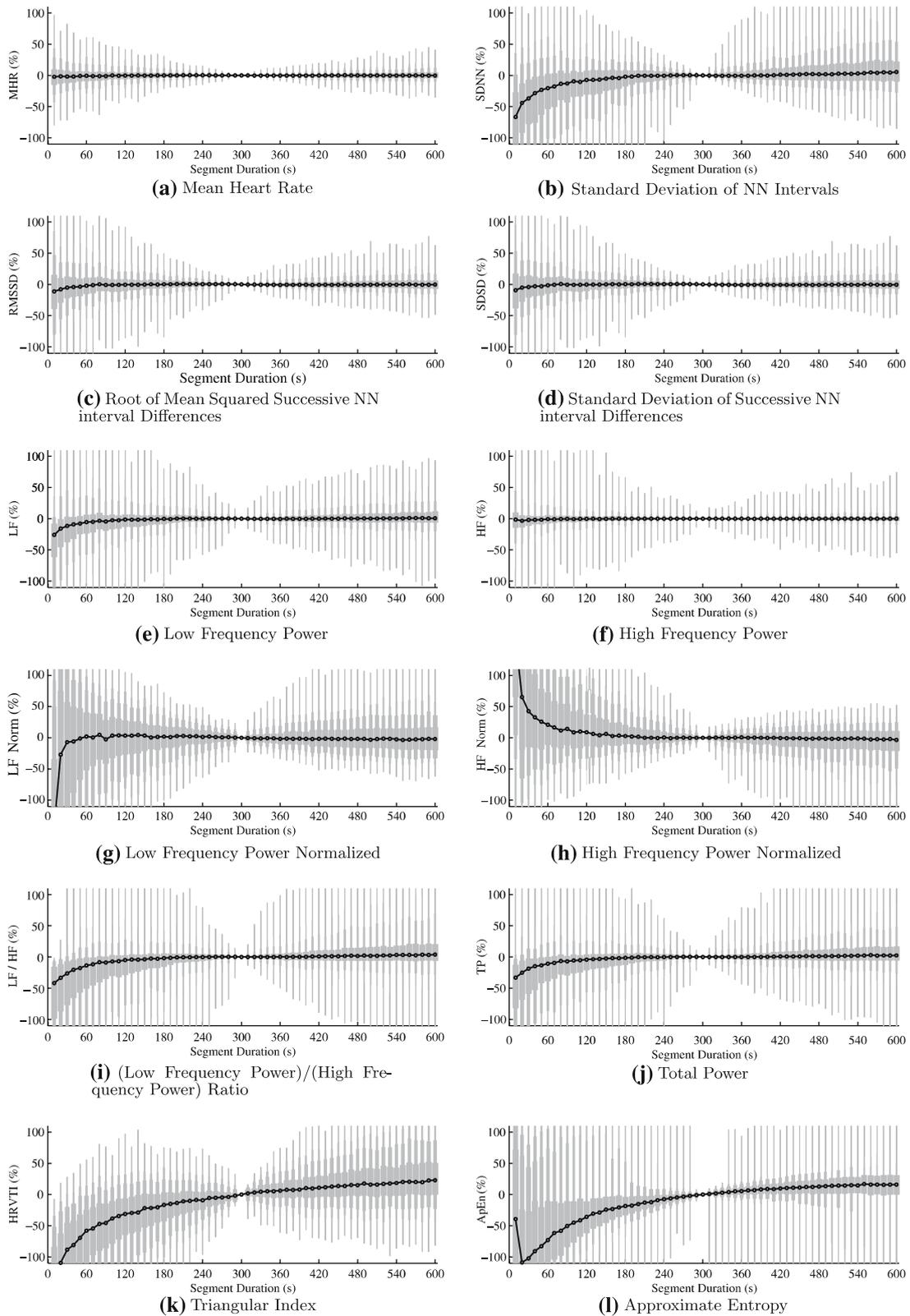
Our results complement those of several other research groups that have studied the reproducibility of HRV metrics in a variety of situations. In general, the objective of previous works has been to assess the reproducibility of a subset of the HRV metrics used to

study a specific patient population or condition of interest. These studies include research on the short- and long-term reproducibility of autonomic measures in supine and standing positions [31], reproducibility of HRV responses to graded lower body negative pressure [20, 34], reliability of short-term HRV measures during exercise [3, 28, 30, 62], reproducibility of measures of cardiovascular autonomic nervous function in middle age and elderly subjects [23], reproducibility of frequency domain HRV metrics before and after a standardized meal [18], HRV analysis reproducibility in the chronic phase of myocardial infarction [5], reproducibility of HRV from short-term recording during manoeuvres in normal subjects [9], stability of short recordings in time [54] and reproducibility of HRV metrics obtained from short-term sampling records [21, 37]. In addition to the reproducibility of 24 h and 5 min records, assessing the reproducibility of HRV metrics calculated from very short records has significant practical importance [25], since it is not always possible to obtain 5 min records due to instrumentation constraints or study design [17]. For example, many 12-lead ECG instruments acquire record of only 10 s in duration.

The main limitation of this work is that we analyzed data from ambulatory recordings, possibly not collected under stationary conditions as recommended by the task force. This may partly explain the lower reliability of the normal sinus rhythm database metrics. These subjects may also have been more active than the other groups.

## 5 Conclusion

We estimated the reliability of 11 metrics calculated from 5 min records and compared the accuracy of the HRV metrics calculated from ECG records spanning 10 s to 10 min to that calculated from a 5 min record and proposed a new methodology and visualization technique which can be used to assess the segment length required for a given degree of accuracy. Our results indicate that all the HRV metrics are sensitive to changes in the duration of the ECG segment. Most of the total HRV is caused by drift and nonstationary effects that cannot be estimated accurately from short segments. HF, SDDSD, and RMSSD had the best overall performance because they both effectively highpass filter the NN interval series and are thereby less sensitive to these effects than the other metrics. The results of the reliability study demonstrate that the MHR was much more reproducible than any of the other HRV metrics.



**Fig. 1** Plots of the difference between eleven different heart rate metrics calculated from a range of segment durations and a 5 min baseline segment. This figure shows the mean heart rate and the other eleven plots show HRV metrics. Each plot shows the median difference and three ranges calculated from 696 15-

min records. The differences are shown as a percentage of the standard deviation of metric calculated from the 5 min baseline of all 696 records. The gray vertical bars show four different ranges of each metric: 25–75, 10–90, 5–95, 1–99%

**Table 1** Summary of a reliability analysis of the cardiac arrhythmia suppression trail (CAST) database

Metric	$\hat{\sigma}$	$\hat{\sigma}_\mu$	2.5%	$\hat{\rho}$ (%)	97.5%
MHR (Hz)	0.001	0.054	97.6	98.0	98.3
SDNN (ms)	0.123	0.281	64.5	69.5	73.8
HRVTI	3.419	9.250	68.5	73.0	76.9
RMSSD (ms)	0.020	0.188	88.8	90.6	92.1
SDSD (ms)	0.020	0.189	88.8	90.6	92.1
LF (ms <sup>2</sup> )	0.088	0.236	68.3	72.8	76.7
HF (ms <sup>2</sup> )	0.005	0.047	87.8	89.7	91.3
LF Norm	0.006	0.023	76.8	80.2	83.2
HF Norm	0.004	0.014	72.3	76.3	79.8
LF/HF	0.000	0.000	30.8	38.7	46.0
TP (ms <sup>2</sup> )	1.758	1.994	46.4	53.1	59.3
ApEn	0.052	0.071	51.5	57.8	63.4

**Table 2** Summary of a reliability analysis of the European ST-T Database

Metric	$\hat{\sigma}$	$\hat{\sigma}_\mu$	2.5%	$\hat{\rho}$ (%)	97.5%
MHR (Hz)	0.001	0.061	98.1	99.0	99.4
SDNN (ms)	0.087	0.624	78.4	87.8	93.3
HRVTI	2.403	11.418	69.8	82.6	90.3
RMSSD (ms)	0.021	0.210	83.5	90.8	95.0
SDSD (ms)	0.021	0.211	83.5	90.8	95.0
LF (ms <sup>2</sup> )	0.342	1.150	61.1	77.1	87.0
HF (ms <sup>2</sup> )	0.019	0.069	63.1	78.3	87.8
LF Norm	0.006	0.024	64.7	79.4	88.4
HF Norm	0.003	0.016	71.5	83.7	90.9
LF/HF	0.000	0.000	74.2	85.3	91.9
TP (ms <sup>2</sup> )	2.680	13.066	70.4	83.0	90.5
ApEn	0.046	0.082	41.8	63.9	78.9

**Table 3** Summary of a reliability analysis of the normal sinus rhythm RR interval database

Metric	$\hat{\sigma}$	$\hat{\sigma}_\mu$	2.5%	$\hat{\rho}$ (%)	97.5%
MHR (Hz)	0.002	0.031	87.3	92.5	95.7
SDNN (ms)	0.204	0.199	25.4	49.4	67.7
HRVTI	4.076	4.915	32.0	54.7	71.4
RMSSD (ms)	0.056	0.135	53.5	70.6	82.1
SDSD (ms)	0.056	0.135	53.5	70.6	82.1
LF (ms <sup>2</sup> )	0.151	0.231	39.5	60.4	75.4
HF (ms <sup>2</sup> )	0.013	0.067	72.5	83.4	90.2
LF Norm	0.007	0.024	63.7	77.6	86.6
HF Norm	0.005	0.017	62.2	76.6	86.0
LF/HF	0.000	0.000	0.0	14.0	39.9
TP (ms <sup>2</sup> )	2.879	1.833	12.9	38.9	60.0
ApEn	0.099	0.039	0.6	28.0	51.5

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